



mquantreg: An R package for estimating generalized linear M-quantile regression models

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General information about the **mquantreg** package

- ▶ Package **mquantreg** is still under development
- ▶ Regression package for the estimation of linear and generalized linear M-quantile regression models
 - ▶ Continuous data
 - ▶ Binary data
 - ▶ Count data (poisson, quasi-poisson, negative binomial)
- ▶ Hypothesis testing via Wald- and Likelihood-Ratio Tests
- ▶ Estimation of q-scores (variable selection; further use in SAE-methods)
- ▶ Model diagnostics using diagnostic plots

Short excursion into theory (1)

What is M-estimation?

- ▶ Huber (1964) formalised a approach to robust estimation of location parameters called M-estimation:

$$\hat{\theta} = \arg \min_{\theta} \left(n^{-1} \sum_{i=1}^n \rho(x_i; \theta) \right)$$

or if $\rho(\cdot)$ is differentiable and convex a solution is easier found with **influence function** $\psi(x; \theta)$:

$$n^{-1} \sum_{i=1}^n \psi(x_i; \hat{\theta}) = 0, \text{ where } \psi(x; \theta) = \frac{\partial}{\partial \theta} \rho(x; \theta)$$

- ▶ examples: $\rho(x; \bar{x}) = (x - \bar{x})^2$, $\rho(x; \tilde{x}) = |x - \tilde{x}|$

Short excursion into theory (2)

Huber Proposal 2 for $\psi(x; \theta)$:

$$\psi_k(x - \theta) = \begin{cases} x - \theta, & \text{if } |x - \theta| < k \\ k \operatorname{sgn}(x - \theta), & \text{if } |x - \theta| \geq k \end{cases}$$

or alternatively $\psi_k(x - \theta) = \max(-k, \min(x - \theta, k))$ with **tuning constant** k , which has to be chosen beforehand.

- ▶ k determines the robustness of the estimator
- ▶ Huber estimator is equivalent to the mean for $k \rightarrow \infty$ and to the median for $k \rightarrow 0$

Short excursion into theory (3)

- ▶ Generalization of the concept to the regression case (Huber, 1973)
- ▶ Further development by Breckling and Chambers (1988) to M-quantile regression for continuous data:

Influence Function (depending on q and k):

$$\psi_{q,k}(y - MQ_{q,k}) = 2[(1 - q)I_{y \leq MQ_{q,k}} + qI_{y > MQ_{q,k}}]\psi_k(y - MQ_{q,k})$$

Estimating equations (depending on q and k):

$$n^{-1} \sum_{i=1}^n \psi_{q,k}(y_i - \hat{MQ}_{q,k}(x_i))x_i = 0, \text{ where } \hat{MQ}_{q,k}(x_i) = x_i^T \hat{\beta}_{q,k}$$

Short excursion into theory (4)

- ▶ M-quantiles are not scale-equivariant: scale parameter $\sigma_{q,k}$ has to be estimated at the same time as regression coefficients
- ▶ The parameters can be estimated via IRLS or ML using Asymmetric Least Informative (ALI) distribution (Bianchi et al. 2018)
- ▶ Categorical data
 - ▶ Generalizations of estimation approach to robust generalised linear models by Cantoni and Ronchetti (2001)
 - ▶ Binary: Chambers, Salvati and Tzavidis (2016)
 - ▶ Poisson: Tzavidis, Ranalli et al. (2015)
 - ▶ Negative Binomial: Chambers, Dreassi and Salvati (2014)

Short excursion into theory (5)

Q-scores:

- ▶ Application of M-quantiles by Kokic et al. (1997)
- ▶ Can be found by solving for q_i^* , which solves

$$MQ_{q_i^*, k}(x_i) = y_i$$

- ▶ Observations with high q-scores have values high values y_i conditional on x_i
- ▶ Q-scores can be used in a lot of different ways in a M-quantile regression context
 - ▶ Model diagnostics
 - ▶ Variable selection (Dawber, 2017)
 - ▶ Useful for example in Small Area Estimation for the estimation of pseudo-random effects (Chambers and Tzavidis (2006))



Estimation

Example for continuous data using a sample of size 1000 from the German EUSILC PUMD (Public Use Microdata) from 2013:

```
1 fit <- mquantreg(formula = income ~ income_self_empl +  
2 income_empl + old_age_benefits + sex,  
3 q = c(0.1, 0.5, 0.7),  
4 k = 1.345,  
5 data = dat,  
6 method = "continuous",  
7 scale.estimator = "Mad",  
8 cont = "lik",  
8 compute.qscores = FALSE)
```

Print and Summary functions (1)

```
1 > print(fit)
2 Call:
3 mquantreg(formula = income ~ income_self_empl + income_
+empl +old_age_benefits + sex, data = dat, q = c
(0.1, 0.5, 0.7), method = "continuous", scale.
estimator = "Mad", compute.qscores = FALSE, cont =
"lik")
4
5 Coefficients:
6            q   (Intercept) income_self_empl income_empl
7 [1,] 0.1    4.936694     0.0004245352  0.4365187
8 [2,] 0.5   12.418848     0.0005556530  0.5109797
9 [3,] 0.7   15.776415     0.0006699692  0.5382013
10
11 old_age_benefits sexFemale
12 0.4866095      1.523643
13 0.4279946      2.754370
14 0.4199769      3.209267
15 Degrees of freedom: 1000 total; 995 residual
```

Print and Summary functions (2)

```
1 > summary(fit)
2 Estimation for m-quantile q = 0.5 with k = 1.345
3 n = 1000
4
5 Coefficients:
6             coefficients   Std. Error
7 (Intercept)    12.418847514 5.827978e-01
8 income_self_empl 0.000555653 3.709961e-05
9 income_empl     0.510979679 1.179632e-02
10 old_age_benefits 0.427994581 3.231350e-02
11 sexFemale       2.754369615 6.575171e-01
12 Estimator for scale parameter sigma: 9.593204
13
14 Residuals:
15 -27.30647 -6.399448 -0.7536041 1.611043 6.858612
16               102.6775
17
18 Proportion of residuals smaller than 0: 0.539
19 Residuals bounded at +- 12.90286
20 Proportion of Huberised residuals: 17.7 %
21 Pseudo R-squared (Bianchi et al, 2015): 0.1356437
```

Variable Selection

```
1   fit_sel <- stepmq(fit)
2 Starting model:
3 income ~ income_self_empl + income_empl + old_age_
    benefits + sex
4
5                               Df      Corr.
6 <none>                  NA 1.0000000
7 - sex                      1 0.9862613
8 - income_self_empl        1 0.9757906
9 - old_age_benefits       1 0.9072897
10 - income_empl            1 0.7163004
11                               Df      Corr.
12 <none>                  NA 1.0000000
13 - income_self_empl       1 0.9761506
14 - old_age_benefits      1 0.9162316
15 - income_empl            1 0.7169594
16                               Df      Corr.
17 <none>                  NA 1.0000000
18 - old_age_benefits      1 0.9241252
19 - income_empl            1 0.7323867
```

Hypothesis Testing (1)

Wald tests for single coefficients:

```
1 > mqwald(fit, id.beta = c(3,4), each = TRUE, q = 0.5)
2 mqwald(fit, id.beta = c(3,4), each = TRUE, q = 0.5)
3 Call:
4 mqwald(object = fit, id.beta = c(3, 4), q = 0.5, each =
TRUE)
5
6 Wald tests for single coefficients at q-value(s) 0.5
7 H0:[1] 0 0
8
9 test-statistic:
10      income_empl old_age_benefits
11 [1,]    1876.353       175.4319
12
13 degrees of freedom
14 [1] 1 1
15
16 p-values
17      income_empl old_age_benefits
18 [1,]          0            0
```



Hypothesis Testing (2)

Likelihood-Ratio Omnibus-Test:

```
1 > mqLRT(fit)
2 Call:
3 mqLRT(object = fit)
4
5 Estimated coefficients to be restricted to zero in
       reduced model:
6
7 [...]
8
9 Result of Likelihood-Ratio test(s):
10          q-value(s)   Chi^2   df    p-value(s)
11 [1,]      0.1     686.3380   4        0
12 [2,]      0.5     986.5277   4        0
13 [3,]      0.7     812.1397   4        0
14
15 The likelihood-ratio test gives significant results
       at a level alpha of 0.05 for the m-quantile(s)
       0.1, 0.5, 0.7 .
```

Graphical Diagnostics (1)

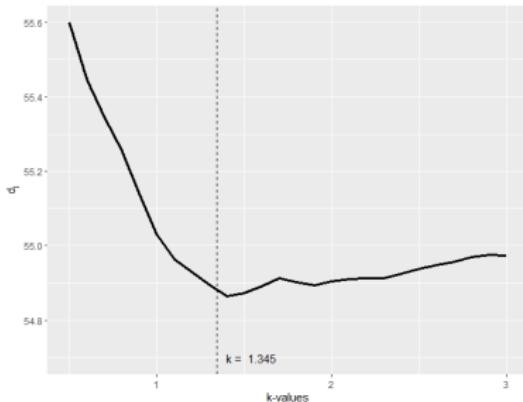
The packages includes the following plotting options (based on Dawber (2017)):

- ▶ "fit.res" - Fitted vs. Residuals Plot
- ▶ "prop.res" - Proportion of Residuals Smaller than 0 vs. Fitted Values
- ▶ "fitted.observed" - Fitted vs. Observed Values
- ▶ "optk" - Comparing the Residual Distribution with Normal Distribution
- ▶ "coef.q" - Coefficients over q
- ▶ "prop.huber.k" - Proportion of Huberised residuals over k
- ▶ "prop.huber.q" - Proportion of Huberised residuals over q

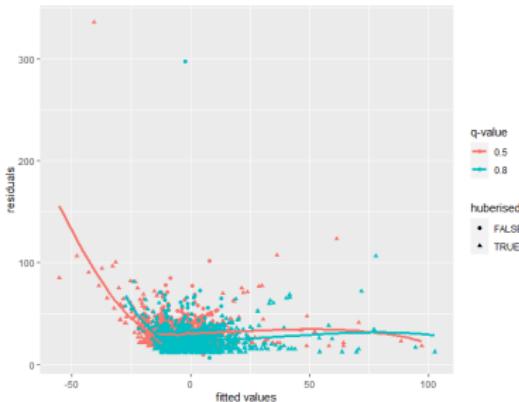
Graphical Diagnostics (2)

```
1 > plot(fit, plottype = "optk")
2 > plot(fit, plottype = "fit.res", add.qvals = c(0.8))
```

Optimal k



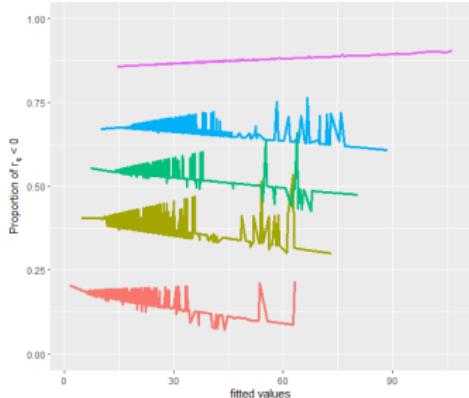
Fitted vs. Residuals



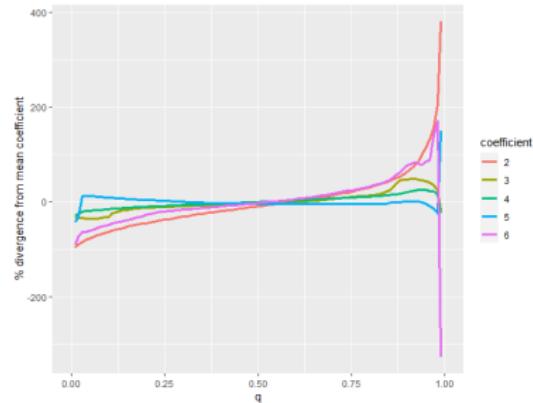
Graphical Diagnostics (3)

```
1 > plot(fit, plottype = "prop.res", add.qvals = c(0.1,0.3,0.7,  
2   0.9))  
> plot(fit, plottype = "coef.q")
```

Fitted vs. Proportion negative Residuals



Coefficients over q-values





Using simulated data under contamination:

```
1 fit_binom <- mquantreg(formula = Y ~ x1 + x2, data =
2   dat, q = 0.5,
3   method = "binom")
```

Summary for simulated example

```
1      Call:  
2      mquantreg(formula = "Y ~ x1 + x2", data = df, q =  
3          0.5, method = "binom")  
4      Estimation for m-quantile q = 0.5 with k = 1.6  
5      n = 1000  
6      Coefficients:  
7          coefficients Std. Error  
8      (Intercept) 0.09966344 0.06557439  
9      x1          0.36810515 0.06782330  
10     x2          0.38303418 0.06855655  
11  
12     Estimator for scale parameter sigma: 0.4971013  
13  
14     Residuals:  
15     -0.7280652 -0.4848358 0.2357902 0.001179976 0.4516844  
16           0.795725  
17  
18     Proportion of residuals smaller than 0: 0.475  
19     Residuals bounded at +- 0.6686012  
20     Proportion of Huberised residuals: 7.9 %  
21     ---  
22     Estimation with Robust Quasi-Likelihood Estimation
```



Conclusion

- ▶ M-quantile regression is a field of continued research
- ▶ These methods are helpful not just as regression models, but can be applied in other contexts as for example SAE
- ▶ Package **mquantreg** can be the foundational package for further use of M-quantile regression
- ▶ In its current state package **mquantreg** already has the functionality users are accustomed to from other regression packages (i.e. `summary()`, `print()`, `plot()` based on classes)



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Thank you for your attention

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